An Essential Primer of Stochastic Processes Su-Chan Park

Assignements

- 1. X_1 and X_2 are independent Gaussian r.v.'s. The mean of X_1 (X_2) is μ_1 (μ_2) and the variance is σ_1^2 (σ_2^2). Show $Y = X_1 + X_2$ is also a Gaussian r.v.. What are the mean and the variance of Y?
- 2. Let X_1, \ldots, X_n be i.i.d. r.v.'s with the probability density function (pdf)

$$P(x) = \frac{1}{\sqrt{2\pi x^3}} \exp\left(-\frac{1}{2x}\right) (x > 0).$$

- (a) Find the generating function $G(\mu) = \langle e^{-\mu x} \rangle \ (\mu > 0).$
- (b) Show that the pdf of an r.v. $Y_n = (X_1 + \dots + X_n)/n^2$ is P(x).
- 3. The Wiener process is defined by the conditional probability

$$P(x,t|y,t') = \frac{1}{\sqrt{4\pi D(t-t')}} \exp\left(-\frac{(x-y)^2}{4D(t-t')}\right),$$

and the Cauchy process by

$$P(x,t|y,t') = \frac{t-t'}{\pi \left[(x-y)^2 + (t-t')^2 \right]},$$

where t > t'.

- (a) Show that both the Wiener process and the Cauchy process satisfy the CK equation.
- (b) By investigating the Lindberg condition, show that the sample paths of the Wiener (Cauchy) process is continuous (discontinuous).
- (c) Calculate W(x|z,t), A(z,t), and B(z,t) for the above two processes.
- 4. In this problem, we will calculate the extinction probability of the Galton-Watson branching process (X_m) with the initial condition $X_0 = 1$.
 - (a) For $p(k) = e^{-1-s}(1+s)^k/k!$, show that the extinction probability for $0 < s \ll 1$ is approximately $\xi = 1-2s$.

- (b) For $p(k) = q(1-q)^k$ (0 < q < 1), find the probability that no individual is left at generation m (that is, find ξ_m defined in the lecture). Find also the generating function $\mathcal{G}_m(z)$.
- 5. Consider a Markov chain with the transition probability

$$T_{mn} = \begin{cases} \frac{n}{N} \frac{(N-n)(1-s)}{n+(N-n)(1-s)} \equiv a_n, & m = n-1, \\ \frac{N-n}{n} \frac{n}{n+(N-n)(1-s)} \equiv c_n, & m = n+1, \\ 1-a_n - c_n, & m = n, \\ 0 & \text{otherwise.} \end{cases}$$

Since $a_0 = c_0 = a_N = c_N = 0$, the states 0 and N are absorbing, so the state space \mathfrak{X} should be finite $[\mathfrak{X} = \{0, 1, 2, \dots, N\}]$. Thus, the system eventually should fall into one of these two absorbing states. By "winning" ("losing"), we mean that the system eventually falls into the n = N (n = 0) absorbing state. Note that this process with s = 0 is martingale in that $\langle n \rangle$ does not change with time. Hence the probability of "winning" can be calculated as

$$\langle n \rangle_0 = N \times \mathbb{P}(\text{"winning"}) + 0 \times \mathbb{P}(\text{"losing"}),$$

 $\Rightarrow \mathbb{P}(\text{"winning"}) = \frac{\langle n \rangle_0}{N}.$

- (a) In case the system starts from n = 1 at t = 0, what is the "winning" probability ?
- (b) Find the expected time to arrive at n = N state for s = 0 provided "winning" will happen (see Prob. 8d).
- 6. Let X_1, \ldots, X_n be i.i.d. r.v.'s with the exponential distribution $\mathbb{P}(X_i > t) = \exp(-\lambda t)$ $(t \ge 0)$. Let $S_n = X_1 + \cdots + X_n$ and N(t) be the number of indices $k \ge 1$ such that $S_k < t$.
 - (a) Show that the pdf for S_n is

$$g_n(t) = \lambda \frac{(\lambda t)^{n-1}}{(n-1)!} e^{-\lambda t}.$$

(b) From $\mathbb{P}(N(t) = n) = \mathbb{P}(S_n < t \text{ and } S_{n+1} > t)$, prove that

$$\mathbb{P}(N(t) = n) = \frac{(\lambda t)^n}{n!} e^{-\lambda t}$$

(c) Show that the solution of the following master equation with $P_n(0) = \delta_{n0}$ is $P_n(t) = \mathbb{P}(N(t) = n)$.

$$\dot{P}_n(t) = -\lambda P_n(t) + \lambda P_{n-1}(t), \dot{P}_0(t) = -\lambda P_0(t)$$

(d) Let t > 0 be fixed, but arbitrary. Show that the element X_k satisfying the condition $S_{k-1} < t \le S_k$ has the pdf

$$v_t(x) = \begin{cases} \lambda^2 x e^{-\lambda x} & \text{for } 0 < x \le t \\ \lambda (1 + \lambda x) e^{-\lambda x} & \text{for } x > t, \end{cases}$$

whose mean approaches $2/\lambda$ as $t \to \infty$.

- (e) Show that the waiting time distribution (for definition, see the lecture note) is $\mathbb{P}(W_t < y) = 1 \exp(-\lambda y)$.
- 7. A master equation on a finite state space $\mathfrak{X} = \{1, \dots, C\}$ satisfies detailed balance iff

$$W_{n_1 n_2} P_{n_2}^e = W_{n_2 n_1} P_{n_1}^e, (1)$$

where $W_{n_1n_2}$ is the transition rate from n_2 to n_1 and P_n^e is the stationary distribution.

(a) It is often necessary to check for the presence or absence of detailed balance in situations where the stationary distribution P_n^e is not explicitly known. To this end, consider a closed loop

$$\mathcal{L} = (n_1 \to n_2 \to n_3 \to \dots \to n_N \to n_1)$$

in the state space along with its 'time-reversed' partner $\bar{\mathcal{L}} = (n_1 \to n_N \to n_{N-1} \to \dots \to n_2 \to n_1)$. Take the product of the transition rates along the loop,

$$\pi(\mathcal{L}) \equiv \prod_{k=1}^{N} W_{n_k n_{k+1}}$$

with $n_{N+1} \equiv n_1$, and show that Eq. (1) holds if and only if $\pi(\mathcal{L}) = \pi(\bar{\mathcal{L}})$ for all possible loops. Note that this requires in particular to construct (at least in principle) the stationary distribution P_n^e from the rates. (b) Make sure that, if the detailed balance is satisfied,

$$W_{n_1n_2} \equiv \left[P_{n_1}^e\right]^{-1/2} W_{n_1n_2} \left[P_{n_2}^e\right]^{1/2}$$

is symmetric and can therefore be diagonalized. Conclude from this that W can be diagonalized and express its eigenfunctions and eigenvalues in those of V.

- 8. In the lecture, the generating function for the continuous time branching process was calculated. In this problem, the r.v. X(t) will be termed as the number of individuals at time t and the system as a population. As in the lecture, the initial number of individuals is m. For simplicity, we set $\lambda = 1$.
 - (a) Find the generating function when $\mu = \lambda = 1$.
 - (b) What is the expected number of individuals at time t?
 - (c) Calculate the extinction probability that the population eventually dies out.
 - (d) Write down the master equation for the survived ensembles, that is, a set of sample paths which never hit0. You should think over the conditional probability

$$\mathbb{P}_x(X(t) = y|S) = \frac{\mathbb{P}_x((X(t) = y) \cap S)}{\mathbb{P}_x(S)},$$

where $\mathbb{P}_x(an \text{ event})$ means the probability that this event occurs provided there were x individuals initially and S means survival by which is meant that extinction will not happen forever. You should recall that this is a Markov process.

- (e) What is the expected number of individuals at time t provided extinction will not happen?
- 9. Prove the following:

$$\underset{n \to \infty}{\text{ms-lim}} \sum_{i=1}^{n} W\left(\alpha t_{i} + (1-\alpha)t_{i-1}\right) \left(W(t_{i}) - W(t_{i-1})\right)$$
$$= \frac{1}{2} \left[W(t)^{2} - W(t_{0})^{2}\right] + \left(\alpha - \frac{1}{2}\right) (t-t_{0}),$$

where W(t) is the Wiener process and $0 \le \alpha \le 1$. One has to prove

$$\left\langle \left(\sum_{i} \Delta W_i^2 - (t - t_0)\right)^2 \right\rangle = 2\sum_{i} (t_i - t_{i-1})^2 \to 0$$

or a similar relation.

10. Show that if a many-variable Itô equation is

$$dx_i = A_i(\boldsymbol{x}, t)dt + B_{ij}(\boldsymbol{x}, t)dW_j,$$

where W_j 's are the independent Wiener processes, then the equivalent FPE is

$$\frac{\partial P}{\partial t} = -\sum_{i} \frac{\partial}{\partial x_{i}} \left[A_{i}(\boldsymbol{x}, t) P \right] + \frac{1}{2} \sum_{ij} \frac{\partial^{2}}{\partial x_{i} \partial x_{j}} \left[\sum_{k} B_{ik} B_{jk} P \right],$$

where $P = P(\boldsymbol{x}, t | \boldsymbol{x}_0, t_0)$. Also prove that if $C_{ij} = B_{ik}S_{kj}$, where S is orthogonal, that is $SS^{T} = 1$, the Itô equation

$$dx_i = A_i(\boldsymbol{x}, t)dt + C_{ij}(\boldsymbol{x}, t)dW_j,$$

has the same FPE as above.

11. Using $\langle \exp(z) \rangle = \exp\left(\frac{1}{2}\langle z^2 \rangle\right)$ for any Gaussian variable z with zero mean, calculate the mean and the autocorrelation function for the geometric Brownian motion (see the lecture note). If the equation for the geometric Brownian motion is interpreted as a Stratonovich one, show that

$$\langle x(t) \rangle = \langle x(0) \rangle \exp\left[\frac{1}{2}c^2(t-t_0)\right],$$

$$\langle x(t)x(s) \rangle = \langle x(0)^2 \rangle \times$$

$$\times \exp\left\{\frac{1}{2}c^2\left[t+s-2t_0+2\min(t-t_0,s-t_0)\right]\right\}.$$

$$(2)$$

12. Consider the following equations (Itô interpretation)

$$d\xi_{\gamma}(t) = -\gamma^{2}\xi_{\gamma}dt + \gamma^{2}dW,$$
$$\frac{dx}{dt} = cx\xi_{\gamma}(t).$$

- (a) Show that $\lim_{\gamma \to \infty} \langle \xi_{\gamma}(t) \xi_{\gamma}(t') \rangle_s = \delta(t t').$
- (b) Show that in the limit of γ → ∞, x(t) is a Gaussian process with mean and autocorrelation function given by Eq. (2).

13. Solve the following SDE (Stratonovich interpretation):

$$(S)dx = \sqrt{x}dW, \quad x(t=0) = 0,$$

and find the generating function $G(k,t) = \langle e^{ikx(t)} \rangle$. Can you conclude that $P(x,t) = \delta(x)$? If not, how do you interpret the result?

14. Project: Considering the following reaction dynamics

$$X \xrightarrow[k_3]{k_1} 0 , \quad X \xrightarrow[k_2]{k_2} 2X ,$$

whose generating function can be obtained exactly, find the moment generating function $G(\mu, t)$ of the following SDE

$$dx = (a + bx)dt + \sqrt{2cx}dW.$$

with a > 0, c > 0, and c > b.

$$G(\mu,t) \equiv \int dx e^{-\mu x} \mathbb{P}(x,t) dx.$$

The initial condition is $P(x, t = 0) = \delta(x - 1)$.